

***Algorithms M2 IF***  
*More on Randomized Algorithms*

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# Review of Probability Theory

# Definition of Probability

To define probabilities we need:

- A “universe” of events  $\Omega$
- A collection of **events**  $\mathcal{E}$  such that for  $E \in \mathcal{E}$  we have  $E \subseteq \Omega$
- A **probability** function  $Pr : \mathcal{E} \rightarrow [0, 1]$

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Example:

- $\Omega = \{1, 2, 3, 4, 5, 6\}$ .
- The following could be events in  $\mathcal{E}$ 
  - $E_3 = \{3\}$
  - $E_{\text{low}} = \{1, 2\}$
  - $E_{\text{odd}} = \{1, 3, 5\}$
- The natural (uniform) probability function would set
  - $Pr[E_3] = 1/6$
  - $Pr[E_{\text{low}}] = 1/3$
  - $Pr[E_{\text{odd}}] = 1/2$

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Example (infinite space):

- $\Omega = [0, 1]$ .
- The following could be events in  $\mathcal{E}$ 
  - $E_3 = \{1/3\}$
  - $E_{\text{low}} = [0, 1/2]$
  - $E_{\text{edge}} = [0, 1/4] \cup [3/4, 1]$
- The natural (uniform) probability function would set
  - $Pr[E_3] = 0$  (why?)
  - $Pr[E_{\text{low}}] = 1/2$
  - $Pr[E_{\text{edge}}] = 1/2$

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A **valid** probability measure satisfies:

- $Pr[\Omega] = 1$
- If  $E_1, E_2, \dots, E_n \in \mathcal{E}$  and for all  $i \neq j$ ,  $E_i \cap E_j = \emptyset$  (mutually disjoint events), then

$$Pr[\cup_{i=1}^n E_i] = \sum_{i=1}^n Pr[E_i]$$

These are called the Kolmogorov probability axioms.

- When  $\Omega$  is finite, the distribution which sets for each  $i \in \Omega$   $Pr[\{i\}] = \frac{1}{|\Omega|}$  is called the **uniform distribution**.

# Probability Basics

Remember: probabilities are **sets** deep down.

- $Pr[\emptyset] = 0$
- If  $E_1 \subseteq E_2$  then  $Pr[E_1] \leq Pr[E_2]$
- $Pr[A \cup B] = Pr[A] + Pr[B] - Pr[A \cap B]$ 
  - Proof?

The last principle can be generalized to give the so-called inclusion-exclusion formula:

$$\begin{aligned} Pr[A_1 \cup A_2 \cup \dots \cup A_n] &= \sum_{i=1}^n Pr[A_i] - \sum_{i_1 \neq i_2=1}^n Pr[A_{i_1} \cap A_{i_2}] + \\ &\quad \sum_{i_1 \neq i_2 \neq i_3=1}^n Pr[A_{i_1} \cap A_{i_2} \cap A_{i_3}] - \dots \end{aligned}$$

# Union bound

A very basic property that follows for **any** collection of events:

$$Pr[\cup_{i=1}^n A_i] \leq \sum_{i=1}^n Pr[A_i]$$

- This is called the **union bound**.
- We often use this bound when  $A_i$  are “bad” events, and we want to show that the probability of one of them happening is small.
  - Main interest: it might be hard to calculate exactly  $Pr[\cup A_i]$ . This allows us to upper bound it without worrying about how each event affects the others.
- The bound becomes an equality only when events are disjoint (mutually exclusive).



# Independence

- Informally: a set of events is **independent**, if knowing that one happened gives us no additional information about the others.
- Formally:  $A, B$  independent if  $Pr[A \cap B] = Pr[A] \cdot Pr[B]$ .
- Formally:  $A_1, \dots, A_n$  independent if for any  $S \subseteq \{1, \dots, n\}$  we have  $Pr[\bigcap_{i \in S} A_i] = \prod_{i \in S} Pr[A_i]$ .

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  - What is the difference between independence for two and for more than two events?
- **Pair-wise independence:**  $A_1, \dots, A_n$  are pair-wise independent iff for any  $i \neq j \in \{1, \dots, n\}$  we have  $Pr[A_i \cap A_j] = Pr[A_i] \cdot Pr[A_j]$ .

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Example: roll a die

- $A$ : result is odd
- $B$ : result is divisible by three
- $C$ : result is  $\geq 4$

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- $A$ : result is odd
- $B$ : result is divisible by three
- $C$ : result is  $\geq 4$
- $A, B$  are independent;  $A, C$  are not;  $B, C$  are independent.

# Conditional Probabilities

- To define independence we asked “Does  $A$  tell us anything about  $B$ ?”
- This corresponds to the notion of **conditional probabilities**:
- Definition:

$$Pr[A | B] = \frac{Pr[A \cap B]}{Pr[B]}$$

- In words: the probability of  $A$ , given  $B$ .
- Note: only makes sense if  $Pr[B] \neq 0$ .

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- So, if  $A, B$  independent, then  $Pr[A | B] = Pr[A]$ .
  - Makes sense!
- Important not to confuse  $Pr[A | B]$  with  $Pr[B | A]$ .
  - $Pr[\text{I sneeze} | \text{I have a cold}] \neq Pr[\text{I have a cold} | \text{I sneeze}]$
- $Pr[A | B]Pr[B] = Pr[B | A]Pr[A] = Pr[A \cap B]$ .

# Useful Tools From Probability Theory



# Expectation

- Random variable: a function  $X : \Omega \rightarrow \mathbb{R}$ .
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Example: we roll a die

- If  $X$  is the number shown,  $X$  is a random variable that takes values in  $\{1, \dots, 6\}$ .
- $Pr[X = 1] = \frac{1}{6}$
- If we roll three dice, let  $Y$  be (a r.v. equal to) their sum
- $Y$  takes values in  $\{3, \dots, 18\}$
- $Pr[Y = 3] = \frac{1}{6^3}$  (why?)

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**Expectation** (discrete variables)

- For a variable  $X : \Omega \rightarrow \mathbb{Z}$  we define

$$E[X] = \sum_{i \in \mathbb{Z}} i \cdot Pr[X = i]$$

- Informally  $E[X]$  is the “average” value of  $X$ .

## Expectation – Geometric distribution

- We have a coin which comes up heads with probability  $p$ . We start flipping it until it comes up heads.
- Let  $X$  be the number of times we flipped it.
- $X$  follows a **geometric distribution**.
- What is  $E[X]$ ?

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$$\begin{aligned} E[X] &= \sum_{i=1}^{\infty} i \Pr[X = i] = \\ &= \sum_{i=1}^{\infty} ip(1-p)^{i-1} = \\ &= -p \sum_{i=0}^{\infty} \frac{d}{dp} ((1-p)^i) = -p \frac{d}{dp} \left( \sum_{i=0}^{\infty} (1-p)^i \right) = \\ &= -p \frac{d}{dp} \left( \frac{1}{p} \right) = \frac{1}{p} \end{aligned}$$

- Makes sense!

# Linearity of Expectations

Why do we like expectations so much?

- Relatively easy to calculate
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- For random variables  $X_1, \dots, X_n$ , constants  $a_1, \dots, a_n \in \mathbb{R}$  we have

$$E\left[\sum_{i=1}^n a_i X_i\right] = \sum_{i=1}^n a_i E[X_i]$$

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- **Important** We don't care if the  $X_i$ 's are independent or not!



## An application: Coupon Collector

- Experiment: we throw a die until we have seen all possible numbers as outcomes.
- Let  $X$  be the number of throws until we stop.
- $E[X] = ?$  (if the die has  $k$  sides)

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- $X_1 = 1$ .
- $X_2$  follows a geom. dist. with probability  $p_2 = \frac{n-1}{n}$ .
- $X_i$  follows a geom. dist. with probability  $p_i = \frac{n-i+1}{n}$ .
- $X = \sum_{i=1}^n X_i$

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$$\begin{aligned} E[X] &= \sum_{i=1}^n E[X_i] = \\ &= \sum_{i=1}^n \frac{1}{p_i} = n \sum_{i=1}^n \frac{1}{n-i+1} \approx n \ln n \end{aligned}$$

## When average is not enough

- Calculating  $E[X]$  is usually only a first step.
- We want to show that  $X$  is “good” (close to  $E[X]$ ) with high probability.
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- **Markov’s inequality**
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- Proof:

$$\begin{aligned} E[X] &= \sum_{i=0}^{\infty} i Pr[X = i] \geq \\ &\sum_{i=\alpha E[X]}^{\infty} i Pr[X = i] \geq \alpha E[X] Pr[X \geq \alpha E[X]] \end{aligned}$$

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- Makes sense!



# Markov collects coupons

Connecting the two previous slides:

- If  $X$  is the number of repetitions until we see all numbers,  
 $E[X] = n \ln n$
- For all  $\alpha > 0$ ,  $Pr[X > \alpha E[X]] \leq \frac{1}{\alpha}$
- $\Rightarrow Pr[X > 100n \ln n] \leq \frac{1}{100}$
- With high probability  $X = O(n \log n)$
- Note: we use the fact that  $X \geq 0$

# Using Variance

# Variance

- A basic way to bound the distance of  $X$  from  $E[X]$  is to calculate  $Var[X]$
- Definition:

$$Var[X] = E[(X - E[X])^2] = E[X^2] - (E[X])^2$$

- Reminder: we often write  $\sigma = \sqrt{Var[X]}$  to denote the **standard deviation** of  $X$ .

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- Reminder: we often write  $\sigma = \sqrt{Var[X]}$  to denote the **standard deviation** of  $X$ .
- Reminder: variance is **not** as nice as expectation.
- Example: in general  $Var[X + Y] \neq Var[X] + Var[Y]$ 
  - However,  $Var[X + Y] = Var[X] + Var[Y]$  if  $X, Y$  independent.

# Using Variance

Chebyshev's inequality:

$$\Pr[|X - E[X]| \geq \alpha] \leq \frac{\text{Var}[X]}{\alpha^2}$$

- In other words, probability that we fall more than  $\alpha\sigma(X)$  away from  $E[X]$  is at most  $\frac{1}{\alpha^2}$ .
- This is why  $\sigma(X) = \sqrt{\text{Var}[X]}$  is called “standard” deviation.

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- Proof:

$$\begin{aligned} \Pr[|X - E[X]| \geq \alpha] &= \Pr[(X - E[X])^2 \geq \alpha^2] \leq \\ &\leq \frac{E[(X - E[X])^2]}{\alpha^2} = \frac{\text{Var}[X]}{\alpha^2} \end{aligned}$$

## Application: Coupon collector again

- Recall Coupon Collector problem:  $X$  is the number of repetitions until we see all outcomes
- $E[X] = n \ln n$
- By Markov,  $Pr[X > 2n \ln n] \leq \frac{1}{2}$

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- Recall that  $X_i$  is repetitions in phase  $i$
- $X = \sum X_i$ , and the  $X_i$ 's are independent
- $Var[X] = \sum Var[X_i]$
- Variance of a geometrically distributed random variable?
  - $Var[Y] = \frac{1-p}{p^2}$ , for  $Y$  geom. with parameter  $p$



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  - $Var[Y] = \frac{1-p}{p^2}$ , for  $Y$  geom. with parameter  $p$

$$\begin{aligned} Var[X] &= \sum_{i=1}^n Var[X_i] \leq \sum_{i=1}^n \left( \frac{n}{n-i+1} \right)^2 = \\ &= n^2 \sum_{i=1}^n \frac{1}{i^2} \leq \frac{\pi^2 n^2}{6} \end{aligned}$$

## Application: Coupon collector again

$$\text{Var}[X] \leq \frac{\pi^2 n^2}{6}$$

We then use Chebyshev's inequality which gives

$$\begin{aligned} \Pr[X > 2n \ln n] &\leq \Pr[|X - n \ln n| > n \ln n] \leq \\ &\leq \frac{n^2 \pi^2 / 6}{(n \ln n)^2} = O\left(\frac{1}{\log^2 n}\right) \end{aligned}$$

Note: Markov's inequality only gives that this probability is at most  $1/2$ .

# Summary

Important lessons to remember.



- Inclusion-Exclusion:  $Pr[A \cup B] = Pr[A] + Pr[B] - Pr[A \cap B]$
- Union bound:  $Pr[A \cup B] \leq Pr[A] + Pr[B]$
- Linearity of Expectation:  $E[X_1 + X_2] = E[X_1] + E[X_2]$
- Markov's inequality:  $Pr[X > a] \leq \frac{E[X]}{a}$
- Variance:  $Var[X] = E[X^2] - E[X]^2$
- Variance only linear for independent variables!
- Chebyshev's inequality:  $Pr[|X - E[X]| > a] \leq \frac{Var[X]}{a^2}$